



1. Introduction

Geostatistical modeling is afflicted by onerous computational effort when the number of locations is vast (the so-called “Big-n” problem).

- Notwithstanding burgeoning literature, spatial inference remains unfeasible for moderate data sets on modest computing environments without access to high-powered architectures.
- Our efforts fall into “meta-” approaches: a massive data set is split into smaller sets, analyzed independently, and local results combined to approximate full Bayesian inference.
- We introduce Bayesian predictive stacking (BPS) in spatial meta-analysis, providing feasible uncertainty quantification without demanding hardware or slowpoke run-time.

2. Conjugate spatial regression - Latent model

Let's consider the following hierarchical model

$$\begin{aligned} y | \omega, \beta, \sigma^2 &\sim \mathbf{N}(X\beta + \omega, \delta^2 \sigma^2 \mathbb{I}_n) \\ \omega | \sigma^2 &\sim \mathbf{N}(0, \sigma^2 \rho_\phi(\mathcal{S}, \mathcal{S})) \\ \beta | \sigma^2 &\sim \mathbf{N}(\mu_\beta, \sigma^2 V_\beta) \\ \sigma^2 &\sim \text{IG}(a_\sigma, b_\sigma). \end{aligned} \quad (1)$$

- $\mathcal{S} = \{s_1, \dots, s_n\} \in \mathbb{R}^2$ be a set of n locations,
- $y = [y(s_i)]^\top$ is $n \times 1$ vector ($i = 1, \dots, n$),
- $X = [x(s_i)^\top]^\top$ is $n \times p$ matrix full rank p ($i = 1, \dots, n$),
- $\mathcal{D} = \{y, X\}$ be the observed dataset over \mathcal{S} .
- $\delta^2 := \tau^2/\sigma^2 \in [0, 1]$ is the noise-to-spatial variance ratio,
- $\omega = [\omega(s_i)]^\top$ is $n \times 1$ latent process ($i = 1, \dots, n$),
- $\rho_\phi(\mathcal{S}, \mathcal{S})$ be the $n \times n$ spatial correlation matrix,
- $\phi \in \mathbb{R}^+$ index spatial correlation function $\rho_\phi(\cdot, \cdot)$.

We cast the model in Equation (1) into an augmented linear system

$$\begin{bmatrix} y \\ \mu_\beta \\ 0 \end{bmatrix}_{y_*} = \begin{bmatrix} X & \mathbb{I}_n \\ \mathbb{I}_p & 0 \\ 0 & \mathbb{I}_n \end{bmatrix}_{X_*} \begin{bmatrix} \beta \\ \omega \\ \gamma \end{bmatrix}_{\gamma_*} + \begin{bmatrix} \eta_1 \\ \eta_2 \\ \eta_3 \end{bmatrix}_{\eta_*}, \quad \eta_* \sim \mathbf{N}(0, \sigma^2 V_*), \quad V_* = \begin{bmatrix} \delta^2 \mathbb{I}_n & 0 & 0 \\ 0 & V_\beta & 0 \\ 0 & 0 & \rho_\phi(\mathcal{S}, \mathcal{S}) \end{bmatrix}.$$

For any fixed $\{\phi, \delta^2\}$ we have conjugacy, leading to the posterior density

$$\begin{aligned} p(\gamma, \sigma^2 | \mathcal{D}) &= p(\sigma^2 | \mathcal{D}) p(\gamma | \sigma^2, \mathcal{D}) \\ &= \text{IG}(\sigma^2 | a_\sigma^*, b_\sigma^*) \mathbf{N}(\gamma | \hat{\gamma}, \sigma^2 M_*). \end{aligned} \quad (2)$$

- $a_\sigma^* = a_\sigma + n/2$,
- $b_\sigma^* = b_\sigma + 1/2 (y_* - X_* \hat{\gamma})^\top V_*^{-1} (y_* - X_* \hat{\gamma})$,
- $M_*^{-1} = X_*^\top V_*^{-1} X_*$,
- $\hat{\gamma} = M_* X_*^\top V_*^{-1} y_*$.

Spatial predictive inference follows in closed form from the posterior distribution (Zhang et al., 2021; Banerjee, 2020)

$$\begin{aligned} p(\omega_u, y_u | \mathcal{D}) &= \int p(y_u | \omega_u, \beta, \sigma^2) p(\omega_u | \omega, \sigma^2) p(\gamma, \sigma^2 | \mathcal{D}) d\gamma d\sigma^2 \\ &= \mathbf{t}_{2a_\sigma^*}(\tilde{\mu}, (b_\sigma^*/a_\sigma^*) \tilde{M}). \end{aligned} \quad (3)$$

- $\mathcal{U} = \{u_1, \dots, u_{n'}\} \in \mathbb{R}^2$ be a set of n' unknown locations,
- $y_u = [y(u_i)]^\top$ is $n' \times 1$ vector ($i = 1, \dots, n'$),
- $X_u = [x(u_i)^\top]^\top$ is $n' \times p$ matrix ($i = 1, \dots, n'$),
- $\omega_u = [\omega(u_i)]^\top$ is $n' \times 1$ latent process ($i = 1, \dots, n'$),
- $\rho_\phi(\mathcal{U}, \mathcal{U})$ be the $n' \times n'$ spatial correlation at \mathcal{U} .
- $\tilde{\mu} = W\hat{\gamma}$, $\tilde{M} = WM_*W^\top + V_e$,
- $V_e = \rho_\phi(\mathcal{U}, \mathcal{U}) - \rho_\phi(\mathcal{U}, \mathcal{S})\rho_\phi^{-1}(\mathcal{S}, \mathcal{S})\rho_\phi^\top(\mathcal{S}, \mathcal{U})$,
- $M_\omega = \rho_\phi(\mathcal{U}, \mathcal{S})\rho_\phi^{-1}(\mathcal{S}, \mathcal{S})$,
- $W = \begin{bmatrix} 0 & M_\omega \\ X_u & M_\omega \end{bmatrix}$, $V_e = \begin{bmatrix} V_\omega & V_\omega \\ V_\omega & V_\omega + \delta^2 \mathbb{I}_{n'} \end{bmatrix}$.

3. Bayesian Stacking of Predictive Densities

Bayesian predictive stacking (BPS) of predictive densities, assimilates different models defining the stacking weights as the following convex optimization problem (Yao et al., 2018; Gneiting and Raftery, 2007):

$$\max_{w \in \mathcal{S}_1^n} \frac{1}{n} \sum_{i=1}^n \log \sum_{j=1}^J w_j p(y_i | \mathcal{D}_{-i}, M_j), \quad (4)$$

- $\mathcal{M} = \{M_1, \dots, M_J\}$ competitive models specified with a collection of values for $\{\delta^2, \phi\}$,
- Equation (4) minimizes KL divergence from the (unknown) true predictive distribution,
- we use the K-fold cross-validation estimate of the expected value of the logarithm score.

4. Accelerated Learning for Spatial Random Fields

In order to accelerate the learning for spatial modeling by BPS

- Data partition:** we divide the full data into K subsets: $\mathcal{D} = \{\mathcal{D}_1, \dots, \mathcal{D}_K\}$.
- Local inference:** Within each partition, we obtain stacked estimations of posterior distributions over the J competitive models

$$\hat{p}(\cdot | \mathcal{D}_k) = \sum_{j=1}^J \hat{z}_{k,j} p(\cdot | \mathcal{D}_k, M_j), \quad (5)$$

$$\text{solving for: } \hat{z}_k = \max_{z_k \in \mathcal{S}_1^{n_k}} \frac{1}{n_k} \sum_{i=1}^{n_k} \log \sum_{j=1}^J z_{k,j} p(y_{k,i} | \mathcal{D}_{k,-j}, M_j).$$

- Global inference:** To achieve inferences on \mathcal{D} , we stack local posterior distribution estimates between partitions

$$\hat{p}(\cdot | \mathcal{D}) = \sum_{k=1}^K \hat{w}_k \hat{p}(\cdot | \mathcal{D}_k), \quad (6)$$

$$\text{solving for: } \hat{w} = \max_{w \in \mathcal{S}_1^K} \frac{1}{n} \sum_{i=1}^n \log \sum_{k=1}^K w_k \sum_{j=1}^J \hat{z}_{k,j} p(y_{k,i} | \mathcal{D}_{k,-j}, M_j)$$

5. Sea Surface Temperature data

Data application focus on Sea Surface Temperature (SST) collected in June 2022

- Trainin data:** $n = 1,000,000$ locations spread all over the world's ocean surface,
- Test data:** $n' = 2,500$ holdout locations to evaluate predictive performances,
- Source:** National Oceanic and Atmospheric Administration (NOAA).

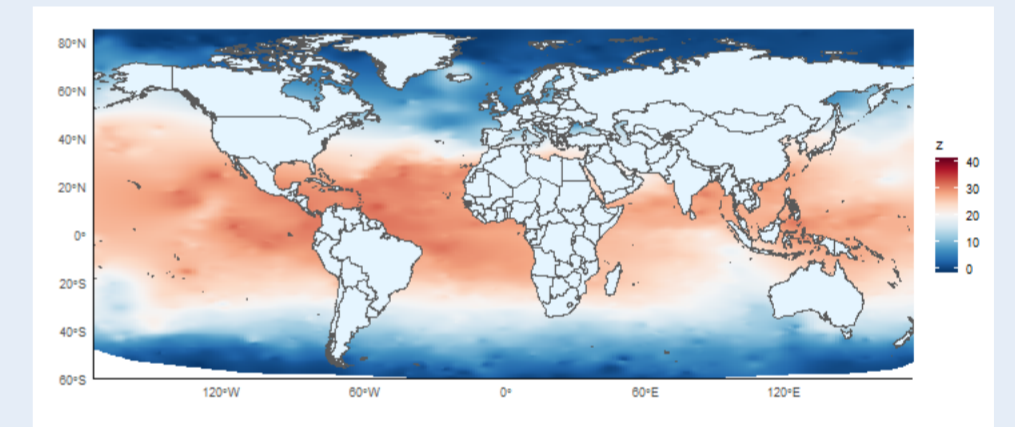


Figure: Holdout data surface interpolation for SST data analysis.

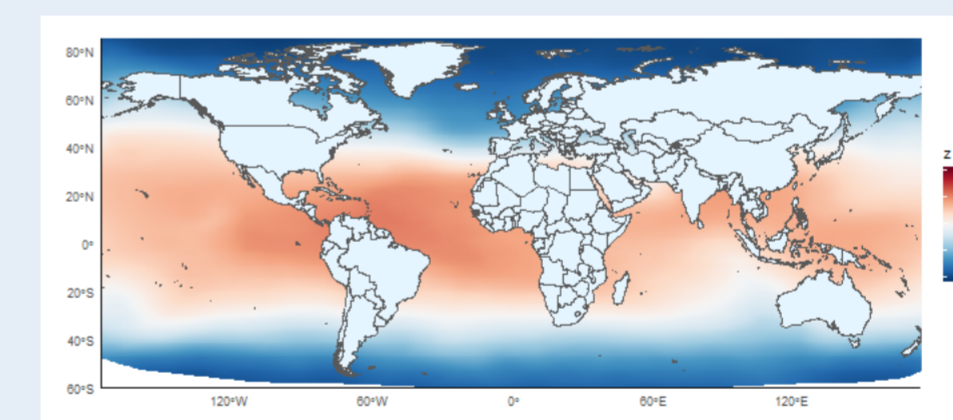


Figure: Predicted (MAP) surface interpolation for SST data analysis.

- Number of subsets:** $K = 2,000$ random partitions (500 locations each),
- Competitive models:** $J = 3$ collections of values for $\{\delta^2, \phi\}$,
- Computing environment:** Intel Core i7-8750H CPU with 5 physical cores.

- Timing:** ≈ 63 minutes (1 hour) for full Bayesian analysis (including model fitting, posterior sampling, and prediction).

- Predictive performances:** RMSPE ≈ 9 times lower w.r.t. Bayesian conjugate model.

	Conjugate linear model	Bayesian predictive stacking
β_0	17.890 (17.87, 17.91)	8.846 (3.382, 12.627)
β_{long}	0.010 (0.01, 0.02)	-0.040 (-0.347, 0.221)
β_{lat}	-0.540 (-0.550, -0.540)	-0.102 (-0.427, 0.412)
σ^2	96.470 (96.21, 96.73)	25.607 (18.032, 38.050)
RMSPE	9.855	1.165
Time	-	63

Table: Sea Surface Temperature data analysis parameter estimates, RMSPE, and computing time in minutes for candidate models. Parameter posterior summary 50 (2.5, 97.5) percentiles.

6. Statistical software

A novel optimized R package was created, named spBPS:

- Introduce the BPS framework for univariate, and multivariate, geostatistical modeling,
- Use Rcpp/C++-based code, allowing faster and more scalable parallel computations,
- Available @luapresicce/spBPS on GitHub (soon on CRAN).

7. References

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